FINDING OF THE EXPECTED INCOME OF THE CLOSED QUEUEING STRUCTURE AND ITS APPLICATION IN TRANSPORT LOGISTICS

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Abstract. Research of the closed exponential queueing structure with one-type messages and the income is conducted. The differential equation in partial derivatives for an income distribution density is received. The ordinary differential equation for its expected income is constructed at particular starting conditions. The offered method of its decision in a case when the intensity of the service of messages, the number of messages in networks, the number of lines of the service in systems, the matrix of probabilities of transitions of messages and the income from transitions between conditions of the closed queueing structure (CQS) depend on time is described. An example when change of parameters has seasonal nature is reviewed. Results of this article can be applied at the prediction of the income of the logistic transport system (LTS).

Introduction

In LTS in practice, the total number of the vehicles moving between various objects and number of loading and unloading crews (workers or, for example, operating loading and unloading tracks at stations) depend on time. Therefore the following CQS can be considered as a LTS model.

Let's consider the closed queueing network consisting of $n+1$ queueing systems $S_0, S_1, \ldots, S_n$, the total number of one-type messages in which in an instant t makes $K(t)$. Usually the system S_0 is understood as environment, and as systems $S_1, S_2, \ldots, S_n$ - concrete queueing systems of network in which the service of messages is made. The queueing networks closed on structure but having not constant total number of served messages, depending on time, are investigated for the first time in special cases in [1] and called the closed queueing structures (CQS). Let $m_i(t)$ - number of service lines in system S_i, $i = \overline{1,n}$, we will put $m_0(t) = K(t)$, and times of messages service of each of service lines are distributed under the exponential law with intensity $\mu_i(t)$, also time-dependent, i, $j = \overline{0,n}$.

Messages for the service get out according to discipline of FIFO. The message, whose service in system S_i ended, with probability $p_{ij}(t)$ passes to system S_j, i, $j = \overline{0,n}$. The matrix of transition probabilities $P(t) = \|p_{ij}(t)\|$, i, $j = \overline{0,n}$, is a matrix of transition probabilities of a nonreducible Markov chain in each moment $0 \leq p_{ij}(t) \leq 1$, $\sum_{j=0}^{n} p_{ij}(t) = 1$. The primal problem of research of the given CQS consists in the asymptotic analysis of the Markov process describing its behavior at a large number of messages.

The condition of structure in a moment t is described by a vector

$$k(t) = (k,t) = (k_1, k_2, ..., k_n, t) = (k_1(t), k_2(t), ..., k_n(t))$$

where $k_i(t)$ - number of messages in system S_i in a moment t, $t \in [0,T]$, $i = \overline{0,n}$, which forms the n-dimensional Markov process with the continuous time and a finite number of conditions. It is apparent that the number of messages in system S_0 is equal to $k_0(t) = K(t) - \sum_{i=1}^{n} k_i(t)$.

Believing that transitions of messages between systems bring in to CQS the particular income, we will set the task of prediction of its total expected income. Let $V(k,t)$ - the complete expected income which will receive CQS in time t if in an initial moment it is in a condition (k,t). It is apparent that $V(k,t) = \sum_{i=0}^{n} V_i(k,t)$, where $V_i(k,t)$ - the expected income which is gained by system S_i in time t if in an initial moment the CQS is in a condition k. In article [2], the concept of distribution density of the expected income of CQS $v(x,t)$, $x = (x_1, x_2, ..., x_n)$ is injected, and the following statement is proved.

Theorem. Income distribution density $v(x,t)$ under a condition that it is differentiated on t and is twice sectionally continuous differentiated on x_i, $i = \overline{1,n}$, satisfies the terms of order of smallness $\varepsilon^2(t) = \dfrac{1}{K^2(t)}$ to the following differential equation in partial derivatives:

$$\frac{\partial v(x,t)}{\partial t} = -\sum_{i=1}^{n} A_i(x,t) \frac{\partial v(x,t)}{\partial x_i} - \frac{\varepsilon(t)}{2} \sum_{i,j=1}^{n} B_{ij}(x,t) \frac{\partial^2 v(x,t)}{\partial x_i \partial x_j} -$$
$$- nK(t)\varepsilon'(t)v(x,t) + r(t) + K(t) \sum_{i,j=0}^{n} \mu_j(t) p_{ji}(t) \min(l_j(t), x_j) r_{ji}(t)$$

(1)

where

$$A_i(x,t) = \sum_{j=1}^{n} \mu_j(t) p_{ji}^*(t) \min(l_j(t), x_j) + \mu_0(t) p_{0i}(t)\left(1 - \sum_{i=1}^{n} x_i\right), \qquad (2)$$

$$p_{ji}^*(t) = \begin{cases} p_{ji}(t) - 1, \ j = i, \\ p_{ji}(t), \ j \neq i, \end{cases} \quad l_j(t) = \frac{m_j(t)}{K(t)}, \ j = \overline{1,n},$$

$B_{ij}(x,t)$ - sectionally continuous function concerning x,

$$B_{ii}(x,t) = \sum_{j=0}^{n} \mu_j(t) q_{ji}^*(t) \min(l_j(t), x_j), \ B_{ij}(x,t) = \mu_i(t) p_{ij}(t) \min(l_i(t), x_i), \quad (3)$$

$$q_{ji}^*(t) = \begin{cases} -1 - p_{ji}(t), \ j = i, \\ -p_{ji}(t), \ j \neq i. \end{cases}$$

1. Differential equation for the average income of structure

Considering (3), expression $\dfrac{\varepsilon(t)}{2} \sum_{i,j=1}^{n} B_{ij}(x,t) \dfrac{\partial^2 v(x,t)}{\partial x_i \partial x_j}$ can be referred to $O(\varepsilon^2(t))$. Therefore we will consider the following equation:

$$\frac{\partial v(k,t)}{\partial t} = -\sum_{i=1}^{n} A_i(x,t) \frac{\partial v(x,t)}{\partial x_i} - nK(t)\varepsilon'(t)v(x,t) +$$

$$+ K(t) \sum_{i,j=0}^{n} \mu_j(t) p_{ji}(t) \min(l_j(t), x_j) r_{ji}(t) + r(t).$$

Having integrated both parts of this equation on $x = (x_1, x_2, ..., x_n)$ in area $G = \left\{ x = (x_1, x_2, ..., x_n) : x_i \geq 0, \ \sum_{i=1}^{n} x_i \leq 1 \right\}$ and having divided both members of equation into the volume of area G, equal to $m(G)$, we receive:

$$\frac{1}{m(G)} \iint \ldots \int_G \frac{\partial v(x,t)}{\partial t} dx = -\frac{1}{m(G)} \sum_{i=1}^{n} \iint \ldots \int_G A_i(x,t) \frac{\partial v(x,t)}{\partial x_i} dx -$$

$$-\frac{nK(t)\varepsilon'(t)}{m(G)} \iint \ldots \int_G v(x,t) dx + \frac{K(t)}{m(G)} \sum_{\substack{i=0, \\ j=1}}^{n} \mu_j(t) p_{ji}(t) r_{ji}(t) \iint \ldots \int_G \min(l_j(t), x_j) dx + \quad (4)$$

$$+ \frac{K(t)}{m(G)} \sum_{i=0}^{n} \mu_0(t) p_{0i}(t) r_{0i}(t) \iint \ldots \int_G \left(1 - \sum_{j=1}^{n} x_j\right) dx + \frac{r(t)}{m(G)} \iint \ldots \int_G dx.$$

Let's consider that in the left part of this equality the change of order of an integration and derivation (we assume that in closed area G function $v(x,t)$ is the continuous) is admissible:

$$\frac{1}{m(G)}\iint\ldots\int_G \frac{\partial v(x,t)}{\partial t}dx = \frac{1}{m(G)}\frac{\partial}{\partial t}\iint\ldots\int_G v(x,t)dx = \frac{d}{dt}\overline{v}_G(t),$$

where $\overline{v}_G(t)$ - an average on x value of the income on condition of change of a reference state (x,t) in area G.

Let's consider integrals in a right member (4). It is apparent that $\frac{nK(t)\varepsilon'(t)}{m(G)}\iint\ldots\int_G v(x,t)dx = nK(t)\varepsilon'(t)\overline{v}_G(t)$. During calculation of the remaining integrals we use integration by parts, and also we will assume that boundary conditions [3] are satisfied:

$$A_i(x,t)v(x,t)\big|_{x\in\Gamma(G)} = 0, \ i = \overline{1,n},$$

were area border G, i.e. $A_n(x,t)v(x,t)\big|_{x_n=0}^{x_n=1-x_1-x_2-\ldots-x_{n-1}} = 0$, $A_{n-1}(x,t)v(x,t)\big|_{x_{n-1}=0}^{x_{n-1}=1-x_1-x_2-\ldots-x_{n-2}-x_n} = 0$, ..., $A_1(x,t)v(x,t)\big|_{x_1=0}^{x_1=1-x_2-x_3-\ldots-x_n} = 0$, which means that the flow of the income through the border of area G is not assumed or that in boundary points of area G reflecting screens are applied. Then, considering that $\frac{\partial A_i(x,t)}{\partial x_i}$ does not depend on x_j, $j = \overline{1,n}$, we receive

$$\frac{1}{m(G)}\iint\ldots\int_G A_i(x,t)\frac{\partial v(x,t)}{\partial x_i}dx = -\frac{\partial A_i(x,t)}{\partial x_i}\overline{v}_G(t), \ i = \overline{1,n}.$$

Therefore we come to the following differential equation

$$\frac{d}{dt}\overline{v}_G(t) = \left[\sum_{i=1}^n \frac{\partial A_i(x,t)}{\partial x_i} - nK(t)\varepsilon'(t)\right]\overline{v}_G(t) +$$

$$+ \frac{K(t)}{m(G)}\sum_{j=1}^n\sum_{i=0}^n \mu_j(t)p_{ji}(t)r_{ji}(t)\iint\ldots\int_G \min(l_j(t),x_j)dx + \qquad (5)$$

$$+ \frac{K(t)}{m(G)}\sum_{i=0}^n \mu_0(t)p_{0i}(t)r_{0i}(t)\iint\ldots\int_G\left(1 - \sum_{j=1}^n x_j\right)dx + \frac{r(t)}{m(G)}\iint\ldots\int_G dx.$$

From (2) we see that the coefficients $A_i(x,t)$ represent piecewise linear functions on x_j, $j = \overline{1,n}$, that is, (5) is a differential equation with a piecewise constant right member. Let's designate a set of indexes of components of a vector

$x = (x_1, x_2, ..., x_n)$ through $\Omega = \{1, 2, ..., n\}$. Let's break Ω into two non-overlapping sets $\Omega_0(\tau)$, $\Omega_1(\tau)$ so that $\Omega_0(\tau) = \{j : l_j(t) < x_j \le 1\}$, $\Omega_1(\tau) = \{j : 0 \le x_j \le l_j(t)\}$, τ – splitting number. At fixed t, the number of splittings of this kind is equal to 2^n, $\tau = \overline{1, 2^n}$. Each splitting will set in a set G not being crossed areas G_τ so that

$$G_\tau = \left\{ x(t) : l_i(t) < x_i \le 1, i \in \Omega_0(\tau); 0 \le x_j \le l_j(t), j \in \Omega_1(\tau); \sum_{i=1}^n x_i \le 1 \right\},$$

$$\tau = 1, 2, ..., 2^n, \quad \bigcup_{\tau=1}^{2^n} G_\tau = G.$$

Now in each area splitting of a phase space we can write down an apparent look (5) and at particular starting conditions find the average expected income for each of the areas G_τ.

Let's set, for example, splitting: $\Omega_0(1) = \{1, 2, ..., n\}$, $\Omega_1(1) = \{\varnothing\}$, $\tau = 1$, that corresponds to existence of queues S_1, S_2, ..., S_n. Then, solving the equation (5) at starting conditions $\overline{v_{G_1}}(0) = S$, it is possible to determine the average expected income changing of a reference state of area $G_1 = \left\{ x(t) : l_i(t) < x_i \le 1, i = \overline{1, n}, \sum_{i=1}^n x_i \le 1 \right\}$. The equation (5) thus looks like

$$\frac{d}{dt} \overline{v_{G_1}}(t) = \left[\mu_0(t) \sum_{i=1}^n p_{0i}(t) - nK(t)\varepsilon'(t) \right] \overline{v_{G_1}}(t) +$$

$$+ \frac{1}{m(G_1)} \left[\sum_{j=1}^n \sum_{i=0}^n \mu_j(t) p_{ji}(t) m_j(t) r_{ji}(t) + \right. \tag{6}$$

$$\left. + K(t) \sum_{i=0}^n \mu_0(t) p_{0i}(t) r_{0i}(t) \underset{G_1}{\iint \ldots \int} \left(1 - \sum_{j=1}^n x_j \right) dx \right] + r(t).$$

2. Example

The transport enterprise (TE, system S_3) at the disposal of which a large number of cars (messages) is available, sends the cars for realization of a number of particular transportations between various cities (environment, system S_0). After that they come back to the TE base, before having passed in two points (systems S_1 and S_2) technical inspection which can also include car repairs. The number of functioning lines of a service in points S_1 and S_2 in a moment t are equal respec-

tively to $m_1(t)$ and $m_2(t)$. In system S_3 loading of cars before the flight in which are engaged $m_3(t)$ loading crews (service lines) is carried out.

A similar situation can arise when cars come back from the environment and are unloaded in two warehouses TE (systems S_1 and S_2); in this case $m_1(t)$ and $m_2(t)$ - number of crews of unloading accordingly in the warehouse S_1 and S_2.

In both cases the model of functioning of transportations is CQS represented in Figure 1.

It is apparent that in this case the matrix $P(t)$ looks like

$$P(t) = \left\| p_{ij}(t) \right\|_{4 \times 4} = \begin{pmatrix} 0 & p_{01}(t) & p_{02}(t) & 0 \\ 0 & 0 & 0 & 1 \\ 0 & 0 & 0 & 1 \\ 1 & 0 & 0 & 0 \end{pmatrix}, \ p_{01}(t) + p_{02}(t) = 1, \ \forall t \in [0,T].$$
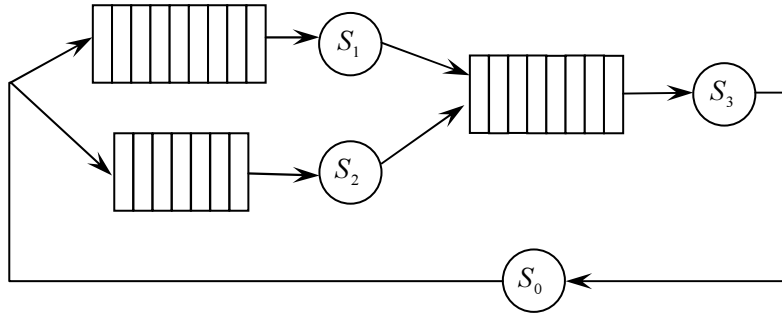

Fig. 1. Model of cars, movement

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Let the change of the CQS parameters have a seasonal nature, for example, they can have one value during the winter period and others in the summer. Let's establish that

$$K(t) = \begin{cases} K_1, t \in \left[0, \dfrac{T}{2}\right), \\ K_2, t \in \left[\dfrac{T}{2}, T\right], \end{cases} \quad \mu_j(t) = \begin{cases} \mu_j^{(1)}, t \in \left[0, \dfrac{T}{2}\right), \\ \mu_j^{(2)}, t \in \left[\dfrac{T}{2}, T\right], \end{cases} \quad j = \overline{0,3},$$

$$m_j(t) = \begin{cases} m_j^{(1)}, t \in \left[0, \dfrac{T}{2}\right), \\ m_j^{(2)}, t \in \left[\dfrac{T}{2}, T\right], \end{cases} \quad j = \overline{1,3}, \quad p_{0i}(t) = \begin{cases} p_{0i}^{(1)}, t \in \left[0, \dfrac{T}{2}\right), \\ p_{0i}^{(2)}, t \in \left[\dfrac{T}{2}, T\right], \end{cases} \quad i = 1,2.$$

Also let the income from transitions between conditions of CQS also be step functions with two intervals of constancy:

$$r_{ji}(t) = \begin{cases} r_{ji}^{(1)}, t \in \left[0, \dfrac{T}{2}\right), \\ r_{ji}^{(2)}, t \in \left[\dfrac{T}{2}, T\right], \end{cases} \quad i,j = \overline{0,3}, \quad r(t) = \begin{cases} r^{(1)}, t \in \left[0, \dfrac{T}{2}\right), \\ r^{(2)}, t \in \left[\dfrac{T}{2}, T\right]. \end{cases}$$

It is clear that $\varepsilon'(t) = 0$.

Let's find area volume $G_1 = \left\{ x : l_i < x_i \leq 1, i \in \{1,2,3\}; 0 \leq x_0 \leq l_0; \sum_{i=1}^{3} x_i \leq 1 \right\}$. We have

$$m(G_1) = \iiint\limits_{G_1} dx = \int\limits_{l_1}^{1-l_2} dx_1 \int\limits_{l_2}^{1-x_1} dx_2 \int\limits_{l_3}^{1-x_1-x_2} dx_3 = \frac{(1-l_1-l_2)^2(1-l_1-l_2-3l_3)}{6}.$$

It is possible also to find integral $\iiint\limits_{G_1}(1-x_1-x_2-x_3)dx$:

$$\iiint\limits_{G_1}(1-x_1-x_2-x_3)dx = \int\limits_{l_1}^{1-l_2} dx_1 \int\limits_{l_2}^{1-x_1} dx_2 \int\limits_{l_3}^{1-x_1-x_2}(1-x_1-x_2-x_3)dx_3 =$$

$$= \frac{(1-l_1-l_2)^2\left((1-l_1-l_2)^2 - 4(1-l_1-l_2)l_3 + 6l_3^2\right)}{24}.$$

Let's first find the expected income of CQS $\overline{v}_1(t)$ on an interval $\left[0, \dfrac{T}{2}\right)$. The equation (6) in this case looks like:

$$\frac{d\overline{v}_1(t)}{dt} = \mu_0^{(1)}\overline{v}_1(t) + \frac{1}{m(G_1)}\Big[\mu_1^{(1)}m_1^{(1)}r_{13}^{(1)} + \mu_2^{(1)}m_2^{(1)}r_{23}^{(1)} + \mu_3^{(1)}m_3^{(1)}r_{30}^{(1)} +$$

$$+ K_1\mu_0^{(1)}\left(p_{01}^{(1)}r_{01}^{(1)} + p_{02}^{(1)}r_{02}^{(1)}\right)\iiint\limits_{G_1}(1-x_1-x_2-x_3)dx + r^{(1)}\Big] = \mu_0^{(1)}\overline{v}_1(t) + r_1.$$

And let $\overline{v}_1(0) = v_0^{(1)}$. Its decision is

$$\overline{v}_1(t) = v_0^{(1)}e^{\mu_0^{(1)}t} + r_1\int_0^t e^{\mu_0^{(1)}(t-\tau)}d\tau = v_0^{(1)}e^{\mu_0^{(1)}t} + \frac{r_1}{\mu_0^{(1)}}\left(e^{\mu_0^{(1)}t} - 1\right) =$$

$$= \left(v_0^{(1)} + \frac{r_1}{\mu_0^{(1)}}\right)e^{\mu_0^{(1)}t} - \frac{r_1}{\mu_0^{(1)}}, \ t \in \left[0, \frac{T}{2}\right].$$

Let's now find the expected income $\overline{v}_2(t)$ of CQS on interval $\left[\dfrac{T}{2}, T\right]$. The equation (6) on this interval becomes

$$\frac{d\overline{v}_2(t)}{dt} = \mu_0^{(2)}\overline{v}_2(t) + r_2, \tag{7}$$

where $r_2 = \dfrac{1}{m(G_1)}\Big[\mu_1^{(2)}m_1^{(2)}r_{13}^{(2)} + \mu_2^{(2)}m_2^{(2)}r_{23}^{(2)} + \mu_3^{(2)}m_3^{(2)}r_{30}^{(2)} + K_2\mu_0^{(2)}\big(p_{01}^{(2)}r_{01}^{(2)} + p_{02}^{(2)}r_{02}^{(2)}\big) \times$

$\times\iiint\limits_{G_1}(1 - x_1 - x_2 - x_3)dx + r^{(2)}\Big]$, as starting conditions it is necessary to take

$\overline{v}_2\left(\dfrac{T}{2}\right) = \overline{v}_1\left(\dfrac{T}{2}\right)$. The solution of the equation (7) is

$$\overline{v}_2(t) = \left[\left(v_0^{(1)} + \frac{r_1}{\mu_0^{(1)}}\right)e^{\mu_0^{(1)}\frac{T}{2}} - \frac{r_1}{\mu_0^{(1)}} + \frac{r_2}{\mu_0^{(2)}}\right]e^{\mu_0^{(2)}t} - \frac{r_2}{\mu_0^{(2)}}, \ t \in \left[\frac{T}{2}, T\right].$$

References

[1] Matalytski M., Rusilko T., Mathematical analysis of stochastic models for claim processing in insurance companies, GrSU, Grodno 2007 (in Russian).
[2] Kiturko O., Matalytski M., Rusilko T., Asymptotic analysis of the total expected income of closed queueing structure with the one-type messages and application, Wiestnik GrUP 2013, 1(2) (in Russian).
[3] Tikhonov V., Mironov V., Markov processes, Sov. Radio, Moscow 1977 (in Russian).